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JACKKNIFE FOR VARIANCE ANALYSIS OF MULTIFACTOR  
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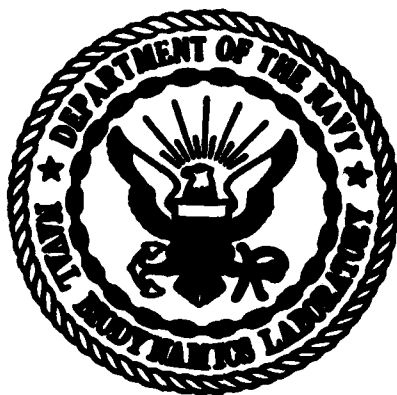
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Jackknife for Variance Analysis of  
Multifactor Experiments

Robert C. Carter  
and  
Alvah C. Bittner, Jr.



May 1982

NAVAL BIODYNAMICS LABORATORY  
New Orleans, Louisiana

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20. ABSTRACT (Continue on reverse side if necessary and identify by block number) A method is presented for analyzing effects of multi-factor experimental treatments on the variance of a dependent variable. The method is based on the statistical jackknife. It enables the analyst to enhance the power of an analysis by using the degrees of freedom associated with the random factor (e.g., subjects, in a behavioral experiment) of a multifactor design. The method is suitable for investigating hypotheses about trends of variances. A computer program is appended which calculates the jackknife variance estimates and other useful statistics.		

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## SUMMARY PAGE

### PROBLEM

Usually an experiment is conducted to show how the experimental treatments affect mean changes of some (dependent) variable. Sometimes there is interest in how the treatments affect the spread of the dependent variable observations.

### FINDINGS

A method is presented for showing how the variability of the dependent variable is affected by treatments. The treatments may be defined in terms of several variables, effects of which can be isolated and evaluated. A FORTRAN language computer program is appended which implements the method.

### RECOMMENDATION

Use the method described herein to show how multifactor treatments affect the variances of observations.

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## Introduction

How was the variability of observations affected by experimental treatments? This question is often of interest during the analysis of an experiment. There are two typical reasons for such interest. First, intertreatment homogeneity of variances is an assumption required for analysis of effects of the treatments on the means of the observations (Winer, 1971). Therefore, a careful analyst assesses the credibility of the hypothesis of homogeneity of variance even though his primary interest is in changes of the means. In addition, sometimes the effect of the treatments on the variances of observations is the analysts' main interest (Bitter & Carter, 1982). This would be the case, for instance, if the treatment were repeated experience with some task (e.g., a mental test) and one wished to investigate the theory (Jensen, 1980) that the variance of observations increases with the amount of experience the subjects have had with the task.

Many methods are available for investigating the hypothesis of homogeneity of variances in experiments having only one independent variable (i.e., treatment) which is used at only a few levels of its range. These methods are summarized by Winer (1971) and Hollander and Wolfe (1975). Such methods are not directly applicable to experiments having multiple independent variables. In addition, they are not applicable for assessing trends in the variances associated with an independent variable whose range is sampled at many levels (e.g., Jensen's theory mentioned above implies an increasing trend in the variances as the independent variable of practice increases).

Games (1978) has suggested an approach to analysis of the variability of observations in multifactor experiments. In essence, he proposed that the natural log transformed ( $\ln$ ) variances be considered as observations in an experimental design having one observation per treatment condition. This ingenious idea, built upon earlier suggestions of Scheffe (1959), is adequate for many purposes. Its primary drawback is that it uses the interaction mean square to estimate its error variance. This use precludes investigations of the interaction. In addition, the strength of a test depends heavily on the number of degrees of freedom in the estimate of the statistical error (i.e., interaction mean square). If the experiment had  $I$  observations at each of  $J$  levels of one treatment by  $K$  levels of another treatment, Games' technique provides an error estimate with degrees of freedom equal to the product:  $(K-1)(J-1)$ . Hence the error estimate can be quite precise if either  $J$  or  $K$  is small. Increasing the number of observations,  $I$ , does not help directly because they are all collapsed into one observation (variance) per treatment condition. It appears that a better technique would somehow take greater advantage of the  $I$  observations per treatment condition. Then the statistical test of the variances would have adequate degrees of freedom if the number of observations were reasonable.

The statistical jackknife (Mosteller & Tukey, 1977) enables the analyst to use the degrees of freedom associated with the observations. It has been used to test homogeneity of variances (Hollander & Wolfe, 1973) in single-treatment experiments. Games, Keselman, and Clinch (1979) endorse its use in multifactor experiments.

### The Jackknife

The statistical jackknife is, as the name implies, a handy general-purpose tool. The purpose for which it is used here is to approximate the distribution of variances in each treatment-condition of an experiment. According to Mosteller and Tukey (1977), the jackknife works well for this purpose as long as the data's sample distribution does not have straggling tails or abrupt end points, is roughly symmetrical, and has a sufficient number of observations.

The jackknife produces numbers which have a distribution like that of a combination of the original data (e.g., the variance is such a combination). If there are  $I$  observations available, then  $I$  variance surrogate numbers can be produced by the jackknife. These numbers can be analyzed to test the hypothesis of homogeneity of variance by using the same procedure used to test the hypothesis of homogeneity of the means of the original observations. The advantage of the unity of analysis made possible by the jackknife is difficult to overemphasize. Obviously, analysis of an experiment is simpler if the same procedure can be used (on the observations and the jackknife variances) to test for effects on the means and variances. Complete parallelism of the two analyses lends clarity of thought, in addition to simplicity of analysis. As an illustration of the type of analysis which can be done, the BMDP computer program P2V (Dixon, 1981) can be used to test multifactor experiments for main effects, trends, and interactions of means (and jackknife variances). To take advantage of this program for inferences on variances only jackknife variance estimates need to be prepared. A computer program will be presented which prepares jackknife variances for analysis.

#### A Computer Program to Prepare Jackknife Variances

The FORTRAN computer program of APPENDIX A produces the jackknife variance surrogate numbers referred to above. A jackknife variance (Hollander & Wolfe, 1977) is produced corresponding to each observation in each treatment condition (which may be defined by multiple treatment factors). Corresponding to observation number  $i$  in treatment  $m$  is

$$S_i^{*2} = N \ln(S_{a11}^2) - (N-1) \ln(S_i^2)$$

where  $S^{*2}$  is a jackknife variance,  $S_{a11}^2$  and  $S_i^2$  are the usual variance estimates calculated, respectively, with all observations in treatment  $m$  or all but observation  $i$  where  $N$  is the number of observations (e.g., subjects) in treatment  $m$ . The natural log-transformation ( $\ln$ ) is employed to reduce the effect of kurtosis on the dispersion of the variance estimates (Scheffe, 1959).

In APPENDIX A,  $S_{a11}^2$  is taken from the diagonal of a variance-covariance matrix of all the observations in all the treatments, and  $S_i^2$  is a diagonal element of the variance-covariance matrix of all observations except the  $i$ th in each treatment condition. The variance-covariance matrix is generated by a subroutine named CORAN (UNIVAC, 1969). The jackknife variances are then punched on computer cards in the same format as the original data. The jackknife variances may then be analyzed conveniently using the same method used for the original data.

In addition, the program in APPENDIX A prints the original data and  $S^2_{all}$  for each treatment condition, and punches on computer cards the correlations among the observations in all pairs of treatment conditions. The original data are printed for identification and verification purposes. The variances,  $S^2_{all}$ , for each treatment condition are printed for the information of the user. Furthermore, the correlations are punched for convenience of analysts who wish to explore the correlation structure of repeated measurements using a method such as that described by Steiger (1980). All of this provides the analyst with options not ordinarily available for more complete description of the data.



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APPENDIX A

A FORTRAN Language Computer Program to Estimate Jackknife Variances

## Variance Jackknife

A-2

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C      THIS IS A FORTRAN LANGUAGE COMPUTER PROGRAM TO CALCULATE JACKKNIFE
C      VARIANCE ESTIMATES AND OTHER USEFUL STATISTICS. THE PROGRAM USES
C      CORAN, A UNIVAC STAT PACK SUBROUTINE TO CALCULATE CORRELATIONS AND
C      VARIANCES. ANOTHER SUBROUTINE COULD BE SUBSTITUTED FOR CORAN.
      DIMENSION DATA(25,15), CORR(15,15), DCORR(15,15), DDATA(15,15),
      *PCOR(15,15), Y(15), SYMR(15), XBAR(15), DDDATA(3,25), (15,5)
      *PCOR(15,15), Y(15), SYMR(15), XBAR(15), DDDATA(3,25), (15,5)
      *PCOR(15,15), Y(15), SYMR(15), XBAR(15), DDDATA(3,25), (15,5)
C      NSUR, NDAY, ARE THE NUMBER OF LEVELS OF TWO INDEPENDENT VARIABLES.
      NSUR=15
      NDAY=15
      NK=1
      NT=NSUR
      NI=NSUR
      NJ=NDAY
C      JSUR, JDAY, IDEN ARE THREE INDEPENDENT VARIABLES.
      DO 100 JSUR=1, NSUR
      DO 100 JDAY=1, NDAY
      DO 100 IDEN=1, 3
C      READ DATA
      READ(5,200) (DDATA(IDEN, JSUR, JDAY)
200  FORMAT (24X, F7.2, 21X F7.2/24X F7.2, 21X, F7.2/24X, F7.2)
100  CONTINUE
      DO 100 IDEN=1, 3
C      WRITE DATA FOR IDENTIFICATION AND VERIFICATION.
      WRITE(6,91)
91  FORMAT(1H , 'DATA:')
      DO 100 KSUR=1, NSUR
      WRITE(6,25) (DATA(KSUR, JDAY), JDAY=1, 15)
25  FORMAT(1H , 15F7.2)
100  CONTINUE
C      WRITE CORRELATION MATRIX WITH LEVELS OF JDAY AS THE INDEX OF THE
C      MATRIX.
      CALL CORAN(DATA, NSUR, NDAY, 1, 0, XBAR, CORR, PCOR, 25, 15)
      DO 100 J=1, NDAY
      WRITE(7,26) (CORR(I, J), J=NDAY, I, -1)
26  FORMAT(1H , 15F5.3)
100  CONTINUE
C      CALCULATE AND WRITE VARIANCE FOR EACH LEVEL OF JDAY WITHIN EACH
C      LEVEL OF IDEN.
      CALL CORAN(DATA, NSUR, NDAY, 0, 0, XBAR, CORR, PCOR, 25, 15)
      WRITE(6,201), CORR(M, M), M=1, NDAY)
201  FORMAT(1H0, 'VARIANCES:', 15F7.2)
      L1=1
      L2=15
      DO 100 JSUR=1, NSUR
      K=0
      DO 100 JSUR=1, NSUR
      IF(JSUR.EQ.JSUR) GOTO 100
      K=K+1
      DO 100 JDAY=1, NDAY
      DDATA(K, JDAY)=DATA(JSUR, JDAY)
102  CONTINUE
100  CONTINUE
      N=NSUR-1
      CALL CORAN(DDATA, N, NDAY, 0, 0, XBAR, DCORR, PCOR, 25, 15)
      DO 100 I=1, NDAY
      L=(JSUR-1)*NDAY+I
      CALCULATE JACKKNIFE VARIANCE ESTIMATES.
      Y(L)=NSUR*ALOG(CORR(I, I))-(NSUR-1)*ALOG(DCORR(I, I))
100  CONTINUE

```

# Variance Jackknife

A-3

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C      WRITE JACKKNIFE VARIANCE ESTIMATES.
      WRITE(7,27)(V(T),I=L1,L2)
27     FORMAT(1H .1D F7.2, / 1H .5F7.2)
      L1=L1+15
      L2=L2+15
301    CONTINUE
302    CONTINUE
      STOP
      END
  
```